Matrix correlations for high-dimensional data: the modified RV-coefficient

A. K. Smilde1,*, H. A. L. Kiers2, S. Bijlsma3, C. M. Rubingh3 and M. J. van Erk3

1Biosystems Data Analysis, Swammerdam Institute for Life Sciences, University of Amsterdam, Nieuwe Achtergracht 166, 1018 WV Amsterdam, 2Heymans Institute, University of Groningen, Groningen and 3TNO Quality of Life, Utrechtseweg 48, 3704 HE Zeist, The Netherlands

Received on July 21, 2008; revised on October 30, 2008; accepted on December 06, 2008

Advance Access publication December 10, 2008

ABSTRACT

Motivation: Modern functional genomics generates high-dimensional datasets. It is often convenient to have a single simple number characterizing the relationship between pairs of such high-dimensional datasets in a comprehensive way. Matrix correlations are such numbers and are appealing since they can be interpreted in the same way as Pearson’s correlations familiar to biologists. The high-dimensionality of functional genomics data is, however, problematic for existing matrix correlations. The motivation of this article is 2-fold: (i) we introduce the idea of matrix correlations to the bioinformatics community and (ii) we give an improvement of the most promising matrix correlation coefficient (the RV-coefficient) circumventing the problems of high-dimensional data.

Results: The modified RV-coefficient can be used in high-dimensional data analysis studies as an easy measure of common information of two datasets. This is shown by theoretical arguments, simulations and applications to two real-life examples from functional genomics, i.e. a transcriptomics and metabolomics example.

Availability: The Matlab m-files of the methods presented can be downloaded from http://www.bdagroup.nl.

Contact: a.k.smilde@uva.nl

1 INTRODUCTION

Functional genomics research generates high-dimensional data, e.g. transcriptomics, proteomics or metabolomics data. The central characteristic of these types of data is the low sample-to-variable ratio. Transcriptomics (or gene-expression) data typically has thousands of variables and the number of samples is in the order of tens to hundred. Similar characteristics hold for proteomics and metabolomics data. Often multiple datasets are available (i.e. measured) on the same samples of the biological system. This calls for data fusion methods: methods that are able to extract the mutual information from all datasets simultaneously (Alter et al., 2003).

A first useful step in such a data fusion strategy is to probe the similarity between pairs of datasets in a simple and comprehensive way (Smilde et al., 2005b). Matrix correlations can be used for this purpose. These correlations take values between zero and one, defining a scale of similarity between two matrices. This scale can be interpreted in much the same way as the absolute value of the Pearson correlation coefficient known to biologists. Hence, its use in functional genomics data fusion can be straightforward.

Matrix correlations have already a long history in multivariate analysis (Robert and Escoufier, 1976; Yanai, 1974). A comprehensive overview is given in Ramsay et al. (1984). For this article, we focus our attention on the RV-coefficient as a typical example of a matrix correlation already in use in metabolomics (Smilde et al., 2005b). While using the RV-coefficient in a transcriptomics study, we ran into problems: the RV-coefficient gave high values in almost all cases. This pointed to trivial results. We explain this trivial result (i.e. the break-down of the RV-coefficient for high-dimensional data) and give a solution to circumvent this unwanted behavior.

2 METHODS

2.1 Matrix correlations

The idea of a matrix correlation is to provide a measure of the similarity of matrices. We start our explanation with matrices X((I×J) and Y((J×J) sharing the row-mode. The latter means that different types of measurements, e.g. transcriptomics and metabolomics, are performed on the same physical samples (the requirement that both matrices have an equal number of columns will be relaxed later). The mapping r:R^I×J→[0,1] is called a correlation function if for all non-zero scalars a and b for X and Y not both zero holds that

\[ C1: r(aX, Y) = r(X, bY) = r(X, Y) \]
\[ C2: r(X, Y) = r(Y, X) \]
\[ C3: r(X, Y) = 1 \text{ if } X = bY \]
\[ C4: r(X, Y) = 0 \text{ if } X'Y' = 0 \]

where 'iff' is the abbreviation of "if and only if" (Ramsay et al., 1984). Matrices can be similar in a variety of ways; this means that rule C3 can be changed, e.g. X and Y can have a correlation of one if they only differ by an orthogonal rotation (X=YQ with Q'Q=I). In that case, the arrangement of the I points (rows) of X and those of Y is essentially equal apart from the rotation. An example of a matrix correlation satisfying C1 to C4 is the absolute value of

\[ r_{tr}(X, Y) = \frac{tr(XY)}{\sqrt{tr(X^2)tr(Y^2)}} \] (2)

which is based on the inner product of two matrices.
A commonly used matrix correlation which allows for a different number of columns in $X(I \times J_1)$ and $Y(I \times J_2)$ is the RV-coefficient (Robert and Escoufier, 1976):

$$RV(X, Y) = \frac{\text{tr}(XX'YY')}{\sqrt{\text{tr}(XX')\text{tr}(YY')}}$$

which is an orientation independent measure, i.e., rotations of the two matrices do not affect the RV-coefficient (it satisfies $C1, C2$ and $C4$, as well as a relaxed version of $C3$; Appendix). This is usually a desirable property since in many functional genomics applications similarities of the configuration of the samples generated by the two matrices is of interest and not their specific orientation. Stated otherwise, the relationships between the samples are of interest not their absolute positions in space.

The RV-coefficient can also be written using the singular value decomposition (SVD) of both $X$ and $Y$:

$$X = U_1D_1V_1' = T_1V_1'$$
$$Y = U_2D_2V_2' = T_2V_2'$$

where $U_1$ and $U_2$ are $I \times I$ orthogonal matrices; $D_1$ and $D_2$ are $I \times I$ diagonal matrices with the singular values of $X$ and $Y$, respectively, on their diagonals; $V_1(J_1 \times I)$ and $V_2(J_2 \times I)$ are column-orthogonal matrices ($V_1V_1' = V_2V_2' = I$). Then it holds that (Ramsay, 1984)

$$RV(X, Y) = r(\text{Vec}(XX'), \text{Vec}(YY'))$$

which can easily be verified by substitution and links the two matrix correlation coefficients $r$ and $RV$. Equation (5) shows that directions in $X$ and $Y$ with more importance (i.e., with high singular values) are given more importance in calculating the RV-coefficients. This property is valuable in high-dimensional data because the interest is usually in communality in the important dimensions of the matrices.

Alternative expressions for the RV are

$$RV(X, Y) = \frac{\text{ssq}(YY)}{\text{ssq}(XX) \times \text{ssq}(YY)}$$

or

$$RV(X, Y) = \frac{\text{Vec}(XX') \text{Vec}(YY')} {\sqrt{\text{Vec}(XX') \text{Vec}(YY') \text{Vec}(YY) \text{Vec}(YY')}}$$

where ssq means sum-of-squares (the sum of squares of all elements of the corresponding matrix) and Vec($X$) is the symbol for the vectorized version of $X$ (see Appendix). The similarity with the Pearson correlation between two vectors $x$ and $y$ becomes clear when writing the latter as

$$r_p(x, y) = \frac{\sum_{i=1}^{n}(x_i - \overline{x})(y_i - \overline{y})}{\sqrt{\sum_{i=1}^{n}(x_i - \overline{x})^2 \sum_{i=1}^{n}(y_i - \overline{y})^2}}$$

(8)

where $x_i$, $y_i$, and $\overline{x}$, $\overline{y}$ are the typical elements and means, respectively, of the vectors $x$, $y$. Rewriting (8) gives

$$r_p(x, y) = \frac{\overline{x} \overline{y}}{\sqrt{\overline{x}^2 \overline{y}^2}}$$

(9)

where $\overline{x}$ is the column-centered version of $x$ and likewise for $\overline{y}$. Since $\text{Vec}(XX')$ and $\text{Vec}(YY')$ in (7) play the same roles as $\overline{x}$ and $\overline{y}$ it is clear that the RV-coefficient can be interpreted as a correlation coefficient. The interpretation of RV as an association measure becomes even more evident when the uncentered correlation is used, or Tucker’s congruence coefficient (Lorenzo-Seva and Ten Berge, 2006)

$$r_T(x, y) = \frac{\overline{x} \overline{y}}{\sqrt{\overline{x}^2 \overline{y}^2}}$$

(10)

which shows that the RV-coefficient bears also similarities with Tucker’s congruence coefficient.

The RV-coefficient is only independent of a rotation or an overall scaling of the matrices ($RV(X, Y) = RV(\alpha XQ_1, \beta YQ_2)$) for non-zero $\alpha$ and $\beta$ and orthogonal $Q_1$ and $Q_2$. All other preprocessing operations are influencing the RV coefficient, e.g. centering has a profound effect similar to the difference between centered and uncentered correlations. Hence, the user has to make a choice regarding the preprocessing and, thus, the metric in which to compare the matrices. Recommendations to this end are available in the literature (Bro and Smilde, 2003; van den Berg et al., 2006).

### 2.2 Problems with the RV-coefficient

While investigating two gene-expression dataset of sizes $5 \times 130$ and $5 \times 113$ from a functional genomics experiment (Kleeman et al., 2007) high RV-coefficients were found (values between 0.5 and 0.99). These could neither directly be understood from the underlying biology nor from independent calculations more extensively investigating the similarities between the two datasets. Hence, a small set of initial simulations was performed where random matrices of the same sizes were generated and RV-coefficients calculated. Despite the fact that the matrices were drawn from random numbers (N(0,1)) the RV-coefficient was always high. Increasing the sample size of the random matrices to 100 samples showed that the RV-coefficient depends on the sample size.

In an extensive set of simulations with random numbers, the samples sizes were systematically increased while the other sizes (130 and 113) remained the same. The result is shown in Figure 1 and shows the problematic behavior. The behavior of the RV-coefficient was also investigated for moderate and strongly unequal sizes of the matrices. Figure 2 shows that the problematic behavior is already visible at much lower numbers of variables (simulations performed similarly as the ones of Fig. 1).

The reason for the unwanted behavior is as follows. According to (7), the RV-coefficient can be written as

$$RV(X, Y) = a'b$$

(11)

with

$$a = \frac{\text{Vec}(XX')}{\text{ssq}(XX')^{1/2}}$$

and

$$b = \frac{\text{Vec}(YY')}{\text{ssq}(YY')^{1/2}}$$

Now, suppose $X$ and $Y$ are fully random matrices, with elements drawn from standard normal distributions. When $J_1$ is large, $XX'$ can be expected to have diagonal elements close to $J_1$ and off-diagonal elements close to 0.

\begin{figure}[h]
\centering
\includegraphics[width=0.5\textwidth]{fig1.png}
\caption{Simulation of RV-coefficients with different numbers of samples. Plotted is the mean, minimum and maximum for 100 repeats each of the RV as a function of the number of samples.}
\end{figure}
Fig. 2. Simulation of RV-coefficient with different numbers of samples and variables. Plotted are the means of the RV-coefficient from 100 repeats.

Specifically, we can write

\[
\begin{align*}
[X^\prime X]_{ij} &= x_i^\prime x_j + \epsilon_{ij}^2 \\
[X^\prime ]_{ij} &= x_i^\prime x_j = \epsilon_{ij}^2, i \neq j,
\end{align*}
\]

where \( x_i^\prime \) denotes the \( i \)-th row of \( X \), the values \( \epsilon_{ij}^2 \) \((i = 1, \ldots, I)\) can be considered as random draws from a distribution with zero mean and (the same) standard deviations \( \sigma_i \). Likewise, the values \( \epsilon_{ij}^2 \) \((i, j = 1, \ldots, I, i \neq j)\) can be considered as random draws from a distribution with zero mean and (the same) standard deviations \( \tau_i \). The reasons for these distributional properties are as follows. Because \( x \) has random elements from a standard normal distribution, \( E(\epsilon x^\prime x) = 0 \), i.e. the expected value of a sum of \( J_1 \) squares of such values, equals \( J_1 \); the variation across realizations of \( x^\prime x_i \) is the same for all \( i \), because the elements of the vectors \( x_i \) are drawn from the same distributions. Likewise, because \( x \) and \( x_j \) have independent random elements, \( E(\epsilon x^\prime x_i) = 0 \), and the variation across realizations of \( \epsilon x^\prime x_i \) is the same for all \( i, j \) because the elements of all vectors \( x_i \) and \( x_j \) are drawn from the same distributions. When \( J_1 \) is large, \( \sigma_i \) and \( \tau_i \) can be expected to be small compared with \( J_1 \). Indeed, when the elements of \( X \) are drawn from standard normal distributions, it holds that \( \sigma_i = \sqrt{2 J_1} \) and \( \tau_i = \sqrt{J_1} \), as follows from general results on stochastic theory for sums and products (Mood et al., 1974).

Using the above distributional results, we can give approximate descriptions of the normalized vectors \( a \) and \( b \). The numerator of \( a \) is the vector with \( I \) values of \( (J_1 + \epsilon_{ij}^2) \) and \( (I - 1) \) values \( \epsilon_{ij}^2 \). Furthermore, we can approximate the denominator in \( a \) as

\[
(sqqXX)^{1/2} \approx (J_1^2 + 1I \sigma_i^2 + (I - 1)\tau_i^2)^{1/2},
\]

which can be explained as follows. The squares of the \( I \) diagonal elements of \( XX^\prime \) sum to \( \sum_{i} (J_1 + \epsilon_{ij}^2)^2 = \sum_{i} J_1 + 2 \sum_{i} \epsilon_{ij}^2 + \sum_{i} (\epsilon_{ij}^2)^2 \). Now using that \( \sum_{i} J_1^2 = IJ_2^2 \), \( \sum_{i} \epsilon_{ij}^2 = 0 \), and that \( \sum_{i} (\epsilon_{ij}^2)^2 \approx I(1 - I)\tau_i^2 \), which completes the approximation of the sum of the squared diagonal values. Furthermore, the sum of squared off-diagonal values is \( \sum_{i} \sum_{i \neq j} (\epsilon_{ij}^2)^2 \approx I(1 - I)\tau_i^2 \), which completes the approximation of the sum of the squared off-diagonal values. Analogously, the numerator of \( b \) is the vector with \( I \) values of \( (J_2 + \epsilon_{ij}^2) \) and \( (I - 1) \) values \( \epsilon_{ij}^2 \), and the denominator in \( b \) can be approximated as

\[
(sqqYY)^{1/2} \approx (J_2^2 + 1I \sigma_i^2 + (I - 1)\tau_i^2)^{1/2},
\]

Now, the RV-coefficient can be approximated as

\[
RV_1(X, Y) = a \cdot b = \frac{(J_1 J_2)}{(J_1^2 + 1I \sigma_i^2 + (I - 1)\tau_i^2)^{1/2}(J_2^2 + 1I \sigma_i^2 + (I - 1)\tau_i^2)^{1/2}}.
\]

\[
RV_1(X, Y) \approx \frac{(J_1 J_2)}{(J_1^2 + (I - 1)\tau_i^2)^{1/2}(J_2^2 + (I - 1)\tau_i^2)^{1/2}}
\]

\[
RV_1(X, Y) \approx \frac{(J_1 J_2)}{(J_1^2 + (I - 1)\tau_i^2)^{1/2}(J_2^2 + (I - 1)\tau_i^2)^{1/2}}
\]

taking into account that in the numerator of \( a \cdot b \) all terms including the random values \( \epsilon_{ij}^2 \) and \( \epsilon_{ij}^2 \) can be expected to roughly cancel. We can further simplify this expression by using the theoretical values for \( \sigma_i \), \( r_i \), \( \sigma_j \) and \( \tau_j \) to obtain

\[
RV_1(X, Y) \approx \frac{(J_1 J_2)}{(J_1^2 + (I - 1)\tau_i^2)^{1/2}(J_2^2 + (I - 1)\tau_i^2)^{1/2}}
\]

Table 1. Comparison of derived and simulated RV-values

<table>
<thead>
<tr>
<th>I</th>
<th>Derived RV-values</th>
<th>Mean simulated RV-values</th>
</tr>
</thead>
<tbody>
<tr>
<td>20</td>
<td>0.852</td>
<td>0.855</td>
</tr>
<tr>
<td>40</td>
<td>0.747</td>
<td>0.748</td>
</tr>
<tr>
<td>60</td>
<td>0.665</td>
<td>0.666</td>
</tr>
<tr>
<td>80</td>
<td>0.599</td>
<td>0.599</td>
</tr>
<tr>
<td>100</td>
<td>0.545</td>
<td>0.545</td>
</tr>
<tr>
<td>200</td>
<td>0.376</td>
<td>0.376</td>
</tr>
<tr>
<td>300</td>
<td>0.287</td>
<td>0.286</td>
</tr>
<tr>
<td>400</td>
<td>0.232</td>
<td>0.232</td>
</tr>
</tbody>
</table>

From (16) and (17) it can be seen that the value of RV for random data matrices depends on \( I \): for small \( I \), the RV is close to 1, whereas, as \( I \) increases the denominator increases and the value approaches zero. The accuracy of these approximations depends on \( I \) but as has been verified in simulations these approximations are typically quite good. Specifically, approximations by the above approach and mean values for RV over 100 random trials yielded the results as given in Table 1 showing that the average values of RV coefficients are very well approximated by the computation of RV according to (17). Apart from rounding errors, there is a close correspondence between the approximations [i.e. (17)] and the means of the simulated RV-values. These results show again that the RV-value is artificially high for small \( I \). Interestingly, the RV-value for the limiting case, i.e. \( I = 1 \), leads to an RV-value of 1, as is easily verified as follows. In the case of \( I = 1 \), the matrices \( XX^\prime \) and \( YY^\prime \) reduce to single numbers, and normalizing these trivially leads to setting these numbers equal to 1, so that the RV-value (i.e. their product) also equals 1.

2.3 The modified RV-coefficient

2.3.1 Definition A solution to the problem of the RV-coefficient presents itself by considering the nature of the problem: the numerator of (17) does not tend to zero for random data matrices and \( J_1 \) and \( J_2 \). This can be traced back to the diagonal of the matrices \( XX^\prime \) and \( YY^\prime \). Indeed, if these diagonal elements are ignored (or, equivalently, set to zero), then the problem disappears since, e.g. \( \text{vec}(XX^\prime) \) would be a vector of random values randomly varying around zero. After using again (7) this would result in an RV-coefficient of nearly zero, as should be the case for the two random matrices. This is then also exactly our proposal for \( RV_2 \), namely instead of using \( XX^\prime \) use \( [XX^\prime - \text{diag}(XX^\prime)] = XX^\prime - \text{diag}(XX^\prime) \) is a matrix containing only the diagonal elements of \( XX^\prime \) on its diagonal, and zero’s elsewhere. Using the analogous definition for \( YY^\prime \) we get

\[
RV_2(X, Y) = \frac{\text{vec}(XX^\prime) \cdot \text{vec}(YY^\prime)}{\sqrt{\text{vec}(XX^\prime)^\prime \cdot \text{vec}(XX^\prime) \cdot \text{vec}(YY^\prime)^\prime \cdot \text{vec}(YY^\prime)}}
\]

Stated otherwise, ignoring the diagonal elements of \( XX^\prime \) and \( YY^\prime \) gives a new vector \( a \) with \( I \) values of 0 and \( (I - 1) \) values \( \epsilon_{ij}^2 \). This solves the problem because the numerator of (17) when using \( RV_2 \) then becomes zero.

2.3.2 Properties The \( RV_2 \) has different properties than the original RV. The most striking one is that \( RV_2 \) can become negative. Suppose for example that

\[
XX^\prime = \begin{pmatrix} 1 & -0.2 \\ -0.2 & 1 \end{pmatrix}
\]
and

\[ YY' = \begin{pmatrix} 1 & 0.2 \\ 0.2 & 1 \end{pmatrix}, \]

(20)

then using (18) gives \( RV(\mathbf{X}, \mathbf{Y}) = -1 \). If instead

\[ X'X = \begin{pmatrix} 1 & 0.2 \\ 0.2 & 1 \end{pmatrix}, \]

(21)

then \( RV(\mathbf{X}, \mathbf{Y}) = 1 \). The interpretation of \( RV = -1 \) is that the association between the rows of \( \mathbf{X} \) is proportional to the association between the rows of \( \mathbf{Y} \) but with a negative sign (equivalent to a negative Pearson correlation).

The \( RV \) depends only on the cross-products \( \mathbf{X}' \mathbf{Y} \) and \( \mathbf{Y}' \mathbf{Y} \), thus the \( RV \) is also orientation independent. The \( RV \) has values in-between \(-1\) and \(1\).

3 EXAMPLES

3.1 Simulated examples

Two simulation examples will be used to illustrate the working of the modified RV-coefficient. The first example addresses the (too) large values of the original RV-coefficient. This example follows closely the gene-expression dataset in which the problem was initially encountered. Two datasets \( \mathbf{X} \) of size \((I \times 130)\) and \( \mathbf{Y} \) of size \((I \times 113)\) were generated 100 times with standard normal distributed numbers. The number of samples was increased from 20 to 400 with steps of 20. For each simulation run, the modified RV-coefficient was calculated. The results are shown in Figure 3.

Fig. 3. Simulation of the modified RV-coefficient \((RV)\) with different numbers of samples \((J_1 = 130, J_2 = 113)\). Plotted is the mean, minimum and maximum for 100 repeats each of the \( RV \) as a function of the number of samples.

The second example shows the working of the RV-coefficients for the case that the amount of overlap between \( \mathbf{X} \) and \( \mathbf{Y} \) gradually increases. Two matrices \( \mathbf{X} \) and \( \mathbf{Y} \) were simulated both of size \((10 \times 100)\) with random numbers drawn from a \( N(0,1) \) distribution and this was repeated 100 times. Gradually, columns of \( \mathbf{Y} \) are exchanged with those of \( \mathbf{X} \) in steps of 10%, 20%, ….

Hence, the amount of overlap between \( \mathbf{X} \) and \( \mathbf{Y} \) increases. Figure 4 shows the original and modified RV-coefficient. Indeed, the modified RV-coefficient gradually increases whereas the original RV-coefficient already has high values from the start.

Fig. 4. Simulation of the original and modified RV-coefficient with different amounts of overlap. Plotted are the mean, minimum and maximum for 100 repeats \( RV \)'s as a function of the amount of overlap.

Table 2. The original and modified RV-values of the gene-expression data (for abbreviations: see text)

<table>
<thead>
<tr>
<th>Case</th>
<th>Original RV</th>
<th>Modified RV</th>
</tr>
</thead>
<tbody>
<tr>
<td>Ct, VA</td>
<td>0.84</td>
<td>0.57</td>
</tr>
<tr>
<td>Low, VA</td>
<td>0.94</td>
<td>0.87</td>
</tr>
<tr>
<td>High, VA</td>
<td>0.91</td>
<td>0.85</td>
</tr>
<tr>
<td>Ct, IC</td>
<td>0.52</td>
<td>0.27</td>
</tr>
<tr>
<td>Low, IC</td>
<td>0.55</td>
<td>0.18</td>
</tr>
<tr>
<td>High, IC</td>
<td>0.83</td>
<td>0.78</td>
</tr>
</tbody>
</table>

Summarizing, Figures 3 and 4 show that the modified-RV coefficient has the desired behavior: (i) on average it equals zero for not related matrices, (ii) for larger sample sizes its variability decreases and (iii) it increases with an increasing amount of overlap.

3.2 Gene-expression example

The gene-expression example is taken from the paper of Kleeman et al. (2007). In short, gene-expression profiles were measured in livers of female ApoE*3L transgenic mice (E3L mice) from three diet groups \((n = 5\) mice per group): control diet (Ct, no cholesterol), low cholesterol diet (Low) and high cholesterol diet (High). Diets were consumed for 10 weeks. RNA from livers was analyzed using Affymetrix whole-genome mouse array MOE430-2.0. Subsets of genes used for matrix correlation were selected based on functional annotation of genes in biological processes cholesterol metabolism \((C, J = 71)\) and inflammation \((I, J = 66)\), vascular development \((V, J = 69)\) and amino-acid metabolism \((A, J = 91)\). Interest focussed on comparison within the different treatment groups, therefore, the RV-coefficients were calculated for the VA blocks and the IC blocks of gene-expressions within each treatment group. All gene-expression values were expressed as deviations from the average Ct group ones without further preprocessing.

Table 2 shows that the modified RV-coefficient is always lower that the original RV-coefficient as it should be. The modified RV-coefficients are more reasonable from a biological perspective. All four selected biological processes were enriched in the selection
of genes differentially expressed in response to cholesterol feeding. From these, the more pronounced response was found on cholesterol metabolism (both in response to low and high cholesterol) and inflammation (high cholesterol) (Kleeman et al., 2007). The dose-dependent gene-expression responses are likely to result in increased correlation between the matrices. Contrary, it is not reasonable to have high correlations between the groups of genes for the control animals (Ct). Hence, the values of the modified RV-coefficient for the control groups are more reasonable than the original RV-coefficients. Also from a statistical point of view, the modified RV-coefficients are better than the original ones. This will be explained for the numbers in the first row of Table 2. Consensus-principal component analysis (CPCA) is an alternative method to probe similarities between matrices (Smilde et al., 2003). Using PCA on the V and A matrices individually gives explained variances of 70.0% and 66.9% for two principal components, respectively. When using CPCA, the two CPCA components explain 57.4% and 64.9% in each block, respectively. The drop in explained variances per matrix (especially for the V block) means that there is some overlap between the matrices but also differences. This agrees nicely with the much lower value of 0.57 instead of 0.84. Note that CPCA is used here only to judge the performance of the modified RV-coefficient. This method does not give an alternative measure to the RV-coefficient but shows qualitatively the same behavior as the modified RV-coefficient supporting the credibility of the latter.

3.3 Metabolomics example

The modified RV-coefficient was applied to a metabolomics dataset (Smilde et al., 2005a). Metabolites were measured in Escherichia coli as a model system. The metabolites were measured using two analytical chemical methods, namely gas-chromatography-mass spectrometry (GC-MS) and liquid-chromatography-mass spectrometry (LC-MS). This generated two datasets with dimensions 28 × 12553 (GC-MS) and 28 × 2532 (LC-MS) which clearly fit into the framework of our modified RV-coefficient. The original RV-coefficient was 0.79 and the modified RV-coefficient was 0.71. Hence, the difference was not large in this case. The CPCA analysis performed in the original publication (Smilde et al., 2005a) showed that both matrices had overlap, but also a substantial non-overlapping part. Although both types of RV-coefficients did not differ much in this case, the example is shown to illustrate that the modified RV-coefficient gives also a reasonable value in this case.

In Smilde et al. (2005a), a truncation was used prior to calculating the RV-value, i.e. the RV-value was calculated using the first principal components of both matrices. Simulations (results not shown) have pointed out that this approach suffers from the same problems as the RV itself and this approach is therefore not recommended.

4 CONCLUSION

It is often convenient to obtain insight into the relationships between blocks of functional genomics data e.g. as a first step in a data fusion strategy. The modified RV-coefficient is a matrix correlation giving such an insight with a single number between −1 and 1. This number can easily be calculated and interpreted in the same way as an ordinary correlation coefficient. The modified RV-coefficient is theoretically motivated and tested with simulations and real data. The results show that this correlation coefficient is reliable and can be used in bioinformatics practice. This coefficient can also easily be combined with permutation testing for assessing significance (Kazi-Aoual et al., 1995) or with bootstrapping to obtain confidence intervals but that is beyond the scope of this article.

Conflict of Interest: none declared.

REFERENCES


A APPENDIX

A.1 Notation

x (vector) bold lowercase

X (matrix) bold uppercase

i = 1, . . . , I object index

j = 1, . . . , J variable index

r = 1, . . . , R principal component index

A.2 RV and orientations

Changing the orientation of the sample conﬁguration of X and Y can be formalized by using arbitrary orthogonal matrices Q1 and Q2 to rotate X and Y, respectively. Upon deﬁning X = XQ1 and Y = YQ2 and observing that the RV-coefﬁcient can be written to depend only on products XX′, (see (7)) it holds that XQ1(XQ1)′ = XQ1Q′1X′ = XX′ due to the orthogonality property of Q1 (and similarly for Y). Hence, RV(X, Y) = RV(XQ1, YQ2).

The I × I matrices XX′ are called conﬁguration matrices and describe the conﬁguration of the I points (rows of X) in their respective row-spaces. The RV-coefﬁcient only measures differences in conﬁgurations (XX′) and not orientations (Qj).

A.3 RV in Vec notation

The equality that tr(A′B) = Vec(A)′Vec(B) can easily been proven by writing both tr(A′B) and Vec(A)′Vec(B) in terms of the elements of the respective matrices. Using this equality in (3) gives (7).